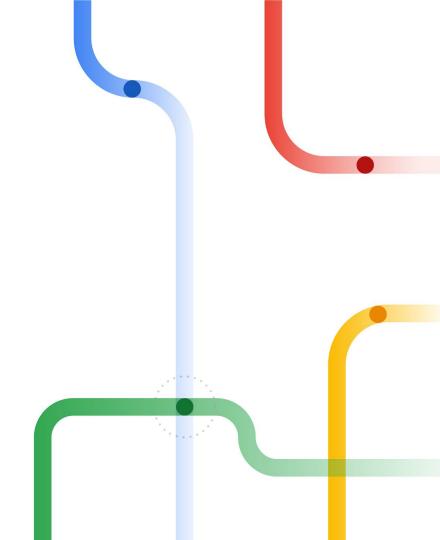
On the benefits of maximum likelihood estimation in regression and forecasting

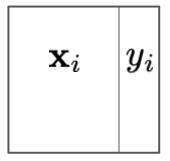
Rajat Sen

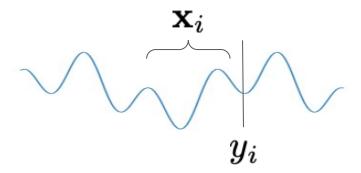
Joint work with Pranjal Awasthi, Abhimanyu Das and Ananda Theertha Suresh

Google Research



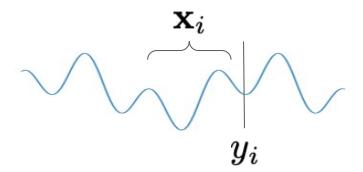
• A training dataset (\mathbf{X}, y^n)





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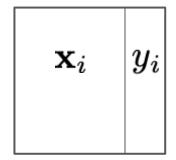
 $egin{array}{c|c} \mathbf{x}_i & y_i \end{array}$

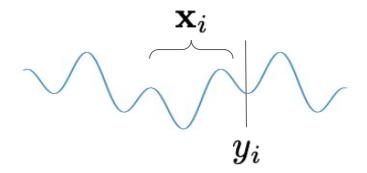


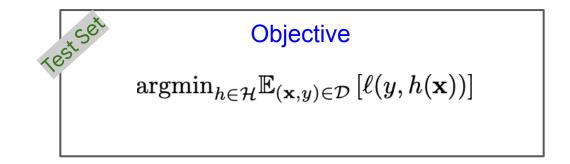
Objective

 $\operatorname{argmin}_{h \in \mathcal{H}} \mathbb{E}_{(\mathbf{x}, y) \in \mathcal{D}} \left[\ell(y, h(\mathbf{x})) \right]$

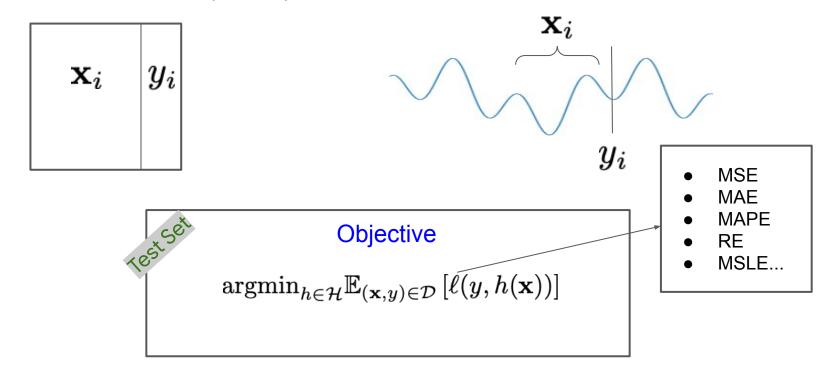
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Approach 1: Target Metric Optimization

• Use target metric as loss and perform Empirical Risk Minimization

$$\hat{h} = \operatorname{argmin}_{h \in \mathcal{H}} \frac{1}{n} \sum_{i=1}^{n} \ell(y_i, h(\mathbf{x}_i))$$

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- Need to train a different model for every target metric
- (PRO) Well understood for many function classes and works reasonably well in practice for many datasets
- (CON) It is provably sub-optimal in some settings like heavy tailed response or covariate [Lugosi et al. 2019]

- Assume that the data is generated from a distribution from a known parametric family of distributions
- Minimize empirical negative log-likelihood over the family of distributions agnostic to the target metric

$$\hat{\boldsymbol{\theta}}_{mle} = \operatorname{argmin}_{\boldsymbol{\theta} \in \Theta} \sum_{i=1}^{n} -\log \left(p(y_i | \mathbf{x}_i; \boldsymbol{\theta}) \right)$$

$$\hat{h}(\boldsymbol{x}) = \operatorname{argmin}_{\hat{y}} \mathbb{E}_{y \sim p(y|\boldsymbol{x}; \hat{\boldsymbol{\theta}}_{mle})} \left[\ell(y, \hat{y}) \right]$$

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- For target metric MSE it is the mean from the MLE
- For MAE it is the median
- Can be computed efficiently for a family of relative errors [Gneiting 11]

$$\left|1-\left(\frac{y}{\hat{y}}\right)^{\beta}\right| o ext{Median of } \sim y^{\beta}p(y|\mathbf{x};\hat{\pmb{ heta}}_{mle})$$

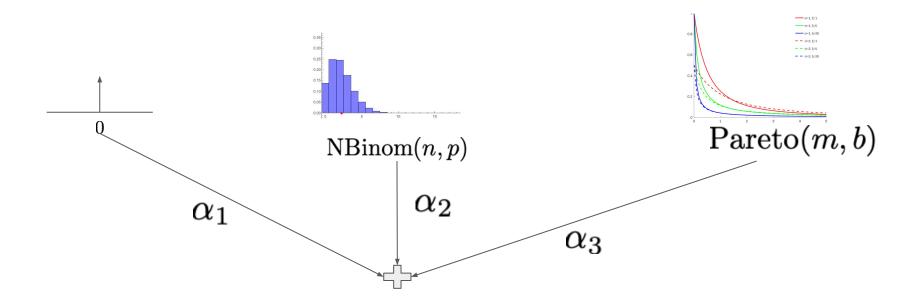
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- (PRO) Only one model can serve many target metrics by adapting the inference
- (PRO) Probabilistic forecasting for free
- (PRO) Can capture inductive biases or domain knowledge by choosing appropriate family like zero inflated negative-binomial for sparse count data
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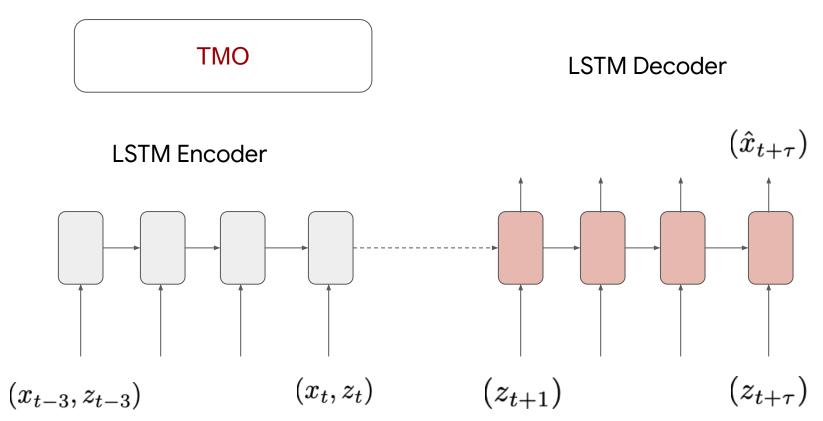
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Choice of Likelihood



Some experimental results....

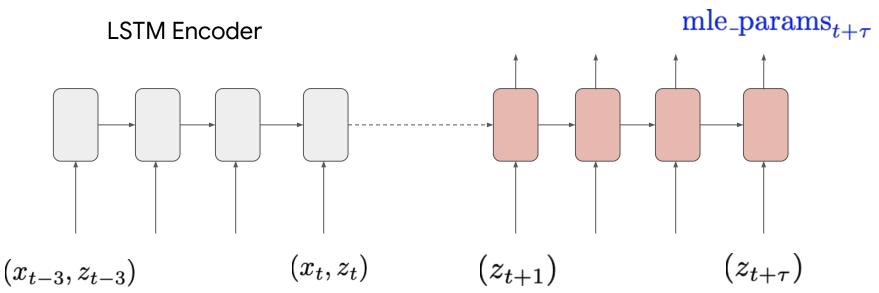
Architecture for Time-Series Models



Architecture for Time-Series Models

MLE with post-hoc Inference

LSTM Decoder



Empirical Results for different Target Metrics

Model	Favorita			M5		
	MAPE	WAPE	RMSE	MAPE	WAPE	RMSE
TMO(MSE)	0.6121±0.0075	0.2891±0.0023	175.3782±0.8235	0.5045±0.004	0.2839±0.0008	7.507±0.023
TMO(MAE)	0.3983 ± 0.0012	0.2258 ± 0.0006	161.4919 ± 0.4748	0.4452 ± 0.0005	0.266 ± 0.0001	7.0503 ± 0.0094
TMO(MAPE)	0.3199 ± 0.0011	0.2528 ± 0.0016	192.3823±1.3871	0.3892 ± 0.0001	0.3143 ± 0.0007	11.3799 ± 0.1965
TMO(Huber)	0.432 ± 0.0033	0.2366 ± 0.0018	164.7006±0.7178	0.4722 ± 0.0007	0.269 ± 0.0002	7.093 ± 0.0133
MLE (ZNBP)	0.3139 ± 0.0011	0.2238 ± 0.0009	$164.6521\!\pm\!1.5185$	$0.3864 {\pm} 0.0001$	0.2677 ± 0.0002	$7.2133 {\pm} 0.0152$

Regression Datasets

Model	Bicycle Share			Gas Turbine		
	MAPE	WAPE	RMSE	MAPE	WAPE	RMSE
TMO(MSE)	0.2503±0.0008	0.1421±0.0003	878.5815±1.3059	0.8884±0.0118	0.3496±0.0041	1.5628±0.0071
TMO(MAE)	0.2594 ± 0.0011	0.1436 ± 0.0003	901.1357±1.4943	0.774 ± 0.0054	0.3389 ± 0.0019	1.5789 ± 0.0067
TMO(MAPE)	0.2382 ± 0.0012	0.1469 ± 0.0012	899.9163±4.8219	0.8108 ± 0.0009	0.8189 ± 0.001	3.0573±0.0019
TMO(Huber)	0.2536 ± 0.0011	0.1414 ± 0.0004	889.1173±1.9654	0.902 ± 0.0128	0.3598 ± 0.0049	1.5992 ± 0.0082
MLE (ZNBP)	0.1969 ± 0.0018	0.1235 ± 0.001	767.4368 ±7.1274	0.9877 ± 0.0019	0.3379 ± 0.0004	1.4547 ± 0.0054

Model	p10QL	p90QL	
TMO (Quantile)	0.0973±0.0002	0.0628±0.0019	
MLE (ZNBP)	0.0788±0.0008	0.0536±0.0007	

In Theory: Can MLE be better than TMO? Is MLE competitive with any estimator?

MLE is competitive with any estimator

- We prove a general result that shows the MLE estimator is competitive with any estimator for a target metric in finite samples under some assumptions
- The competitive ratio depends on the size of the family of distribution (or the size of its cover). This is different from the competitive result in Acharya et al 2016. Also does not follow from classical MVUE results.
- We show that our assumptions are valid for a class of convex likelihood families for least square regression

Application: Poisson regression with Identity Link

We consider fixed design Poisson regression with the identity link function

$$p(y_i = k | \mathbf{x}_i; \boldsymbol{\theta}^*) = \frac{\mu_i^k e^{-\mu_i}}{k!}$$
 where $\mu_i = \langle \boldsymbol{\theta}^*, \mathbf{x}_i \rangle$

We will consider the square loss as our target metric for simplicity

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 We find a (complicated) estimator based on [Lugosi and Mendelson, 19] with guarantee

$$\mathcal{E}(\boldsymbol{\theta}_{\mathrm{est}}) \leq c \cdot \|\boldsymbol{\theta}^*\|^2 \cdot \lambda_{\max}(\Sigma) \left(\frac{d + \log(\frac{1}{\delta})}{n}\right).$$

• Our result allows us to show that the simple MLE based estimator is competitive with the above upto log(n)/n factors

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Gábor Lugosi and Shahar Mendelson. Mean estimation and regression under heavy-taile Computational Mathematics, 19(5):1145–1190, 2019a.

Better than TMO guarantee

Application: Pareto Regression

- We consider the Pareto regression setting where the scale parameter is relevant for the regression
- By virtue of our results we show that MLE based estimator is competitive with a complicated heavy tailed regression estimator [Hsu and Sabato, 16]
- This shows that we are better than TMO which will fail to have sub-gaussian type guarantees for heavy tailed data

Thank you! Questions?